# Non-predation natural mortality of Norway pout (Trisopterus esmarkii) in the North Sea 

Henrik Sparholt, Lena I. Larsen, and J. Rasmus Nielsen


#### Abstract

Sparholt, H., Larsen, L. I., and Nielsen, J. R. 2002. Non-predation natural mortality of Norway pout (Trisopterus esmarkii) in the North Sea. - ICES Journal of Marine Science, 59: 1276-1284.

Based on age disaggregated data on catch rates in bottom trawl surveys, commercial catches, and the number consumed by the North Sea piscivorous predators, new estimates of non-predation natural mortality, M1, are obtained for Norway pout [Trisopterus esmarkii (Nilsson)]. Simple log catch ratio analysis and rough maximum likelihood procedures are applied. The analysis focus on the year classes 1977-1981 and 1987-1991, which are represented in the extensive stomach sampling of North Sea piscivorous fish in 1981 and 1991. Although the M1 of Norway pout varied between the two periods, in both periods it increased with age and was very high for age 2 and older fish $(0.10$ for age $1,1.74$ for age $2,2.58$ for age 3 and 3.05 for age 4 for the 1977-1981 year classes and 0.10 for age 1, 2.03 for age $2,3.04$ for age 3 and 4.39 for age 4 for the 1987-1991 year classes). This difference between the two periods is not significant. Survey data from each quarter of the year show that the main mortality takes place between the 1 st and the 2 nd quarter of the year, i.e. from before to after spawning, thus pointing at spawning as the main factor.


(C) 2002 International Council for the Exploration of the Sea. Published by Elsevier Science Ltd. All rights reserved.
Keywords: commercial catch, natural mortality, North Sea, Norway pout, numbers consumed, population dynamics, trawl survey, Trisopterus esmarkii.

Received 9 October 2001; accepted 6 June 2002.
H. Sparholt and L. I. Larsen: ICES, Palcagade 2-4, DK-1261 Copenhagen K, Denmark; tel: +45 331542 25; e-mail: henriks@ices.dk; lena@ices.dk.J. R. Nielsen: Danish Institute for Fisheries Research, Charlottenlund Castle, DK-2990 Charlottenlund, Denmark; tel: +45 9633 00; e-mail: rn@dfu.min.dk.

## Introduction

Norway pout [Trisopterus esmarkii (Nilsson)] is a small gadoid fish which, in the North Sea, is exposed to a significant fishery for fishmeal and fish oil. The mean annual catch in 1974-1999 was 278000 t (ICES, 2001). It is a short-lived species, which rarely gets older than 5 years. Norway pout is an important prey for cod (Gadus morhua), whiting (Merlangius merlangus), saithe (Pollachius virens) and mackerel (Scomber scombrus) in the North Sea (ICES, 1997), and it is included in the ICES North Sea Multi-Species model (MSVPA) (see e.g. Sparre, 1991). Predation mortality (M2) is estimated by the MSVPA based on extensive stomach analysis of the fish predators, their stock numbers and their consumption rates. For the MSVPA model to run, assumptions about other (non-predation) natural mortality (M1) have to be made. This mortality can be caused by disease, spawning stress, growth stress (Ursin, 1963), etc.

As for most species in the MSVPA model, M1 is assumed to be 0.2 for all ages of Norway pout.

There are, however, serious uncertainties about appropriate M1 values for Norway pout. Ursin (1963), Raitt (1968), and Bailey and Kunzlic (1984) have indicated that natural mortality (M) of Norway pout increases with age. Ursin (1963) suggested that the increase could result from: ". . . the irreversibility of growth in length in a seasonally changing environment, causing a marked loss of condition in winter". Raitt (1968) and Bailey and Kunzlic (1984) observed that total mortality, Z, estimated from survey data increases with age. Bailey and Kunzlic (1984) speculated that this could be due to spawning stress.

In the routine stock assessment made by the ICES Working Group on the Assessment of Demersal Stocks in the North Sea and Skagerrak (e.g. ICES, 2001), natural mortality of Norway pout is assumed to be constant over ages. Based on results of the MSVPA

Table 1. Survey indices $\left(\mathrm{N}^{-1}\right)$ for Norway pout by age.

| Year/age | IBTS 1st quarter |  |  |  |  | EGFS August |  |  |  | SGFS August |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1 | 2 | 3 | 4 | 5 | 0 | 1 | 2 | 3 | 0 | 1 | 2 | 3 |
| 1974 | 11921 | 8100 | 776.0 | 1.0 | 0.1 | - | - | - | - | - | - | - | - |
| 1975 | 4827 | 1808 | 20.7 | 11.0 | 0.5 | - | - | - | - | - | - | - | - |
| 1976 | 4066 | 303 | 13.8 | 0.0 | 0.0 | - | - | - | - | - | - | - | - |
| 1977 | 6095 | 256 | 53.8 | 0.0 | 0.0 | - | - | - | - | - | - | - | - |
| 1978 | 1480 | 551 | 46.9 | 0.3 | 0.0 | - | - | - | - | - | - | - | - |
| 1979 | 2558 | 307 | 73.4 | 0.0 | 0.0 | - | - | - | - | - | - | - | - |
| 1980 | 3275 | 552 | 29.1 | 4.1 | 0.0 | - | - | - | - | - | 1928 | 346 | 12 |
| 1981 | 1092 | 377 | 14.9 | 0.2 | 0.0 | - | - | - | - | - | 185 | 127 | 9 |
| 1982 | 4437 | 256 | 57.3 | 1.2 | 0.0 | 6594 | 2609 | 39 | 77.0 | 8 | 991 | 44 | 22 |
| 1983 | 2326 | 628 | 8.0 | 3.4 | 0.1 | 6067 | 1558 | 114 | 0.4 | 13 | 490 | 91 | 1 |
| 1984 | 3957 | 812 | 56.3 | 0.9 | 0.2 | 457 | 3605 | 359 | 14.0 | 2 | 615 | 69 | 9 |
| 1985 | 2117 | 1423 | 72.9 | 3.2 | 0.0 | 362 | 1201 | 307 | 0.0 | 5 | 636 | 173 | 5 |
| 1986 | 2051 | 385 | 20.0 | 1.1 | 0.0 | 285 | 717 | 150 | 80.0 | 38 | 389 | 54 | 9 |
| 1987 | 3173 | 485 | 62.9 | 3.2 | 0.3 | 8 | 552 | 122 | 0.9 | 7 | 338 | 23 | 1 |
| 1988 | 124 | 724 | 12.9 | 2.1 | 0.6 | 165 | 102 | 134 | 21.0 | 14 | 38 | 209 | 4 |
| 1989 | 2015 | 252 | 171.6 | 2.7 | 0.2 | 1530 | 1274 | 621 | 20.0 | 2 | 382 | 21 | 14 |
| 1990 | 1271 | 758 | 40.4 | 2.7 | 0.0 | 2692 | 917 | 158 | 23.0 | 58 | 206 | 51 | 2 |
| 1991 | 2497 | 677 | 129.0 | 0.0 | 0.0 | 1509 | 683 | 399 | 6.0 | 10 | 732 | 42 | 6 |
| 1992 | 5121 | 902 | 33.0 | 4.7 | 0.0 | 2885 | 6193 | 1069 | 157.0 | 12 | 1715 | 221 | 24 |
| 1993 | 2681 | 2644 | 259.0 | 6.0 | 0.0 | - | - | - | - | 2 | 580 | 329 | 20 |
| 1994 | 1868 | 375 | 67.0 | 2.9 | 0.2 | - | - | - | - | 136 | 387 | 106 | 6 |
| 1995 | 5941 | 785 | 77.0 | 8.6 | 0.0 | - | - | - | - | 37 | 2438 | 234 | 21 |
| 1996 | 912 | 2635 | 234.0 | 4.7 | 0.1 | - | - | - | - | 127 | 412 | 321 | 8 |
| 1997 | 9752 | 1474 | 670.0 | 2.9 | 0.0 | - | - | - | - | 1 | 2154 | 130 | 32 |
| 1998 | 1006 | 5343 | 300.0 | 73.8 | 0.1 | - | - | - | - | - | - | - | - |
| 1999 | 3527 | 597 | 667.0 | 4.0 | 0.3 | - | - | - | - | - | - | - | - |

model, natural mortality is set to 1.6 annually, for all age groups.

This paper tries to resolve the problem of appropriate M1 values by analysing, (1) catch rate indices from research vessel trawl surveys, (2) catch in commercial fisheries and (3) the number of Norway pout consumed by the predators in the North Sea. All data series are disaggregated by age.

## Materials

Standard abundance indices of Norway pout (Table 1) from the International Bottom Trawl Survey (IBTS) in the North Sea, Skagerrak and Kattegat (see e.g. ICES, 1999) in February each year are used to calculate total mortality (Z). Beginning in 1974 the IBTS spatial coverage of the Norway pout distribution area is good and fairly constant over time (Anon., 1999). The IBTS covers mainly depths less than 200 m in the North Sea and less than 250 m in the Skagerrak. According to Poulsen (1968) very few Norway pout are caught at depths greater than 200 m in the North Sea and Skagerrak on shrimp trawl surveys. Albert (1994) found Norway pout at depths greater than 200 m , but very few at depths greater than 300 m .

The IBTS index expresses the mean catch ( $\mathrm{N} \mathrm{h}^{-1}$ ) within a standard area (ICES, 1999). The area comprises 93 statistical rectangles ( $1^{\circ}$ Longitude $\times 0.5^{\circ}$ Latitude) and covers most of the geographical distribution of the Norway pout stock in the North Sea and Skagerrak. The area with depths less than 300 m not covered by the index area is about $15 \%$ of the index area, and the index area is thus assumed to cover the main part of the stock distribution. Between 274 and 527 trawl hauls have been made annually and since 1978 almost all fish were caught using the GOV trawl (Anon., 1999). The data used in the present analyses are taken from the ICES IBTS database.

Data from Scottish and English ground fish surveys were also used. The Scottish Ground Fish Survey (SGFS) is conducted in the third quarter of the year and covers the period 1980-2001. Scotland changed from the Aberdeen trawl to the IBTS GOV trawl in 1998 and therefore no data from 1998 and onwards have been included in order to avoid intercalibration problems. The English Ground Fish Survey (EGFS) is also conducted in the third quarter of the year and covers the period 1982-2001. England changed from a Granton trawl used until 1992, to the IBTS GOV trawl. Again to avoid intercalibration problems no data from later than 1992 were considered. The number of statistical

Table 2. Commercial catch in numbers ('000) at age and in weight ('000 t).

| Year | Age |  |  |  |  | Weight |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 0 | 1 | 2 | 3 | 4 |  |
| 1974 | 6565750 | 25647060 | 821700 | 174830 | 690 | 735.8 |
| 1975 | 10856250 | 20091580 | 2919240 | 16630 | 640 | 559.7 |
| 1976 | 6183000 | 21035390 | 2143130 | 166770 | 510 | 435.4 |
| 1977 | 1715690 | 19868160 | 2413650 | 91210 | 2510 | 389.9 |
| 1978 | 1529650 | 7896880 | 3129160 | 322520 | 3930 | 270.1 |
| 1979 | 1886960 | 15189410 | 2182570 | 264710 | 4120 | 329.2 |
| 1980 | 684950 | 19839730 | 4363080 | 115360 | 7210 | 482.7 |
| 1981 | 37003370 | 5705490 | 3589540 | 174730 | 8080 | 238.5 |
| 1982 | 1245270 | 18654330 | 1202010 | 310030 | 0 | 395.3 |
| 1983 | 3117380 | 16154400 | 4485510 | 48760 | 2120 | 451.4 |
| 1984 | 2232100 | 13793570 | 4956070 | 420160 | 0 | 393.0 |
| 1985 | 684420 | 7512340 | 2475540 | 224400 | 714 | 205.1 |
| 1986 | 5571530 | 3633200 | 1440140 | 81120 | 3010 | 178.4 |
| 1987 | 234700 | 7540110 | 865010 | 17430 | 1010 | 149.3 |
| 1988 | 3886670 | 1158680 | 1428740 | 20100 | 0 | 109.5 |
| 1989 | 5012250 | 5827040 | 540410 | 29630 | 0 | 172.5 |
| 1990 | 1013000 | 5772000 | 1457000 | 49000 | 10000 | 151.6 |
| 1991 | 4219620 | 4704170 | 2141420 | 152290 | 6000 | 192.9 |
| 1992 | 1833230 | 11319480 | 2035240 | 140710 | 3000 | 299.8 |
| 1993 | 1271130 | 4952680 | 2528550 | 93660 | 170 | 184.2 |
| 1994 | 4885090 | 4524270 | 1430790 | 156000 | 0 | 182.3 |
| 1995 | 2392040 | 11789850 | 602180 | 43330 | 0 | 241.0 |
| 1996 | 3241550 | 2788630 | 2308020 | 88640 | 0 | 166.2 |
| 1997 | 451120 | 5782810 | 1035400 | 337390 | 0 | 169.7 |
| 1998 | 432520 | 1404180 | 1538760 | 83050 | 28400 | 79.8 |

rectangles covered at the SGFS is 86 and at the EGFS 73. Both surveys cover the entire North Sea, but Skagerrak is not covered. Normally one haul is made in each rectangle (Cotter, 2001). The data are taken from ICES (2000).

Commercial catch (numbers-at-age per year) from 1974-1998 were taken from ICES routine assessment work (ICES, 2000) (Table 2).

The number of Norway pout consumed by the piscivorous predators in the North Sea (Table 3) was calculated from the MSVPA model based on the Key-run of ICES (1997). The basis for the calculation is the stomach content of the predators, the numbers of the predators and the prey, and the consumption rates of the predators. The stock numbers are integrated parts of the MSVPA calculations and consequently use the assumed M1 values. These are, however, what we in the present paper try to estimate. To avoid circularity in the logic, we estimated the numbers of Norway pout consumed in the years where predator stomachs were sampled (1981 and 1991). For these years the estimates of numbers predated are basically a simple multiplication of predator numbers, consumption rates and fraction of Norway pout in the stomach.

Table 3. Number ('000) of Norway pout consumed by age according to the MSVPA.

|  | Age |  |  |  |
| :--- | ---: | ---: | ---: | ---: |
| Year | 0 | 1 | 2 | 3 |
| 1974 | 199606469 | 94982799 | 1972306 | 212529 |
| 1975 | 353421137 | 75619550 | 5435707 | 11867 |
| 1976 | 189202594 | 116344116 | 5391641 | 197327 |
| 1977 | 87953305 | 76936061 | 7665342 | 80035 |
| 1978 | 128218549 | 38356893 | 6813697 | 414319 |
| 1979 | 158546416 | 58334339 | 3027496 | 217896 |
| 1980 | 131129669 | 75898765 | 5963956 | 59692 |
| 1981 | 228789151 | 35741200 | 10129841 | 219985 |
| 1982 | 209840717 | 86164145 | 1907051 | 261172 |
| 1983 | 172921338 | 65912278 | 7645556 | 48961 |
| 1984 | 208592270 | 52908393 | 7721441 | 292568 |
| 1985 | 143569361 | 46539874 | 5568713 | 127141 |
| 1986 | 233472637 | 30435654 | 2124384 | 93138 |
| 1987 | 207134563 | 37614512 | 1709526 | 50981 |
| 1988 | 221679199 | 12590380 | 3386310 | 4573 |
| 1989 | 196586439 | 30914165 | 1279065 | 75583 |
| 1990 | 250028184 | 29143276 | 3672968 | 40490 |
| 1991 | 238651609 | 29841369 | 2730734 | 236719 |
| 1992 | 177709628 | 46854308 | 2864390 | 64680 |
| 1993 | 236576148 | 28083447 | 6528422 | 146458 |
| 1994 | 261622328 | 24921411 | 2771157 | 189176 |
| 1995 | 223786325 | 58949185 | 2777046 | 87225 |

In order to calculate stock biomass from stock numbers at 1 January mean values of weight at age in quarter 1 and in quarter 4 the previous year were used. Data were taken from ICES (2000).

## Methods

Two methods were used to estimate M1. First, individual time series of data, i.e. the survey data, the commercial catch data and the number consumed, were used to calculate total mortality by Baranov (1918)'s catch equation. Secondly, a simple model is used to estimate fishing mortality (F), other mortality (M1), and predation mortality (M2) by maximum likelihood fitting of the IBTS survey indices, commercial catches and numbers consumed, to the model.
To avoid possible long-term trends when estimating total mortality by Baranov's catch equation, separate analysis have been performed on the year classes 19771981 represented in the stomach data from 1981, and on the year classes 1987-1991 represented in the stomach data from 1991. Within each of the two periods the year class data have been averaged before analysis. This makes it more difficult to calculate confidence intervals of estimated parameters, but has the advantage of increasing the signal to noise ratio. This approach is especially important for the survey data where individual yearly abundance indices are quite variable with large positive correlation between adjacent age groups in a given year (Cotter, 2001).
While the above procedures give independent estimates of $\mathbf{Z}$ for each time series, the aim in the maximum likelihood model is to estimate M1 for age 2, 3 and 4 as consistent as possible with all three data series; IBTS survey indices, commercial catches and numbers consumed.
The model is a forward calculation procedure, which starts with an input stock number of age $1, \mathrm{~N}_{1}$, and successively calculates $\mathrm{N}_{2}, \mathrm{~N}_{3}$, etc. The parameters, F , M1 and M2 by age are estimated by the model on the basis of formulas relating these parameters to (1) the IBTS indices, (2) commercial catches, and (3) numbers consumed.

Catch at age in numbers (C) is assumed to be log normal distributed with mean
$\mathrm{E}[\mathrm{C}(\mathrm{a}, \mathrm{y})]=\mathrm{avN}(\mathrm{a}, \mathrm{y}) * \mathrm{~F}(\mathrm{a}, \mathrm{y})$,
where $\operatorname{avN}(a, y)$ is the mean number of fish of age a present in year $y$, assuming the Baranov's catch equation.

The numbers consumed (D) are treated in a similar way, with F replaced by M2.

The IBTS indices ( $\mathrm{N} \mathrm{h}^{-1}$ ) are also assumed to be log normal distributed with mean
$\mathrm{E}[\operatorname{IBTS}(\mathrm{a}, \mathrm{y})]=\mathrm{q} * \mathrm{~N}(\mathrm{a}, \mathrm{y})$,
where q is the catchability of the IBTS survey, and $N(a, y)$ is the number of fish at the beginning of year $y$.

For each of the two time periods the five year-classes are merged in order to reduce the noise in the data. This implicitly assumes that F, M1 and M2 at age are constant within each of the two time periods.

It is assumed that M1 on age 1 is 0.1 , and that F and q are constant for age 1 and older fish for each of the two time periods. This assumption is justified by the following points: (1) that the spatial distribution of Norway pout of age 1 does not differ much from that of older fish, (2) that the mesh size in the GOV trawl is 10 mm (knot to knot) with a $100 \%$ selection for Norway pout greater than 9 cm (Anon., 1998), which means that almost all Norway pout of age 1 or older entering the trawl will be caught, and (3) that Norway pout of age 2 and older rarely are longer than 20 cm , i.e. not much bigger than those of age 1 and therefore can be expected to have approximately the same catchability.

The data on number consumed at age 4 and older are unfortunately only available as a part of an age $3+$-group. However, as the stock number of age 4 and older are very low compared to age 3 as judged from all available information, the number of age $3+$ consumed can be assumed to be representing only age 3 . Also due to the lack of information on age 4 fish consumed it is assumed that the ratio between number caught and number consumed is equal to that of age 3 .

The IBTS standard area for Norway pout is $284000 \mathrm{~km}^{2}$. The swept area of a 1 h haul is 4 nm times a trawl wing spread of 20 m , i.e. $0.15 \mathrm{~km}^{2}$. The IBTS catchability is measured as the percentage caught by the trawl in the swept area, assuming that the entire stock is distributed in the standard area.

A maximum likelihood estimates are obtained by minimizing the following function:

$$
\begin{aligned}
& \Sigma(\ln (\mathrm{C})-\ln (\mathrm{E}[\mathrm{C}]))^{2}+\Sigma(\ln (\mathrm{D})- \\
&\ln (\mathrm{E}[\mathrm{D}]))^{2}+\Sigma(\ln (\mathrm{qN})-\ln (\mathrm{E}[\mathrm{IBTS}]))^{2}
\end{aligned}
$$

which is the sum of the squared differences between observed and expected values of the survey catch rates, the commercial catches, and the numbers consumed, all log-transformed. The variances of the log-transformed data are assumed to be constant except for age 4 in the commercial catches and age 5 in the IBTS data, which are assumed to have a low precision and are down-weighted by an arbitrary factor of 10 .

The parameters estimated are $N_{1}, F$ at age 1-4 (constant by age), M1 at age 2-4 (variable by age) and the IBTS catchability (constant by age), in total 6 parameters. M2 can be calculated from F based on the proportion of numbers caught to numbers predated. The observations are (1) the mean IBTS indices for the year classes 1977-1981/1987-1991 at age 1, 2, 3, 4 and 5 (5 observations), (2) the mean commercial catch of the

Table 4. Z estimated by time series. Y.c. = year class. $1 \mathrm{q}=$ first quarter of the year, $3 \mathrm{q}=$ third quarter of the year.

| Data source/Age | 0 | 1 | 2 | 3 | 4 |
| :--- | :---: | :---: | :---: | :---: | :---: |
| IBTS 1q, y.c 1977-1981 | - | 1.92 | 2.55 | 2.92 | 3.39 |
| Commercial catch, y.c. 1977-1981 | - | 1.45 | 2.69 | 4.58 | - |
| Numbers consumed, y.c. 1977-1981 | 0.91 | 2.33 | 3.48 | - | - |
| IBTS 1q, y.c 1987-1991 | - | 0.75 | 2.29 | 3.17 | 4.30 |
| SGFS 3q, y.c. 1987-1991 | - | 1.53 | 2.44 | - | - |
| Commercial catch, y.c. 1987-1991 | - | 1.20 | 2.69 | 4.17 |  |
| Numbers consumed, y.c. 1987-1991 | 2.01 | 2.17 | 3.23 | - | - |
| IBTS 1q, entire period 1974-1999 | - | 1.02 | 2.13 | 3.13 | 3.99 |
| EGFS 3q, entire period 1982-1992 | - | 1.70 | 2.29 | - | - |
| SGFS 3q, entire period 1980-1997 | - | 1.73 | 2.54 | - | - |
| Commercial catch | - | 1.53 | 2.78 | 4.41 | - |
| Numbers consumed | 1.35 | 2.45 | 3.45 | - | - |

same year classes at age 1,2, 3, and 4 (4 observations) and (3) the mean number consumed of the same year classes at age 1, 2, and 3 ( 3 observations), in total 12 observations. The IBTS survey data were the only survey data included in the modelling, because these data cover both time periods and because the IBTS data are expected to be more precise than the EGFS and SFGS data due to better spatial coverage and to a significant higher number of annual trawl hauls. The model was run separately for each of the two time periods.

Confidence intervals of the parameters are obtained by the bootstrap method (see e.g. Haddon, 2001).

## Results

Table 4 shows $Z$ by age, calculated from the various time series. In the surveys, the catch rates of 0 group Norway pout are relatively low, probably due to the fact that 0 group Norway pout in the 3rd quarter of the year are small and some can escape through the meshes. Furthermore, a large fraction of the 0 groups is probably still semi-pelagic at that time and will not be caught by the bottom trawl. Norway pout recruit to the commercial fishery as 0 -groups at the end of the year and the total annual catch of them is therefore low. 1-group Norway pout can be regarded as fully recruited to the fishery already from the start of the year. Therefore, the only time series, which might produce realistic $Z$ values for age 0 fish are the numbers consumed.
Z of 0 groups based on numbers consumed was 0.91 for the year classes 1977-1981 and 2.01 for year classes 1987-1991. For age 1 Z was between 1.45 and 2.33 for the year classes 1977-1981, while Z for the year classes 1987-1991, was between 0.75 and 2.17. $Z$ of age 2 increased by about 1.0 for all data series compared to Z for age 1 . For ages 3 and 4 Z increased further compared
to Z for age 2. The increase varied between 0.37 and 2.01 .

The EGFS and the SGFS gave generally slightly higher Z values for age 1 and age 2 . These surveys are conducted in the 3 rd quarter and thus represent fish half a year older than the IBTS data.

Generally, Z calculated from numbers consumed was higher than for the other time series for age 1 and older fish and Z from the commercial catches slightly higher than Z from the surveys.

Table 5 shows the results of the maximum likelihood model runs for each of the two periods.

The model estimates of M1 is 1.74 for age 2, 2.58 for age 3 and 3.05 for age 4 for the year classes 1977-1981. For the year classes 1987-1991 the comparable estimates are 2.03, 3.04 and 4.19. Total mortality, Z , estimated for the two period matches quite well Z estimated from the IBTS survey (Table 4). Fishing mortality on age 1 and older fish is estimated to be 0.29 in the first and 0.10 in the second period. Predation mortality decreases by increasing age, as well as from the first to the second period. For age 1 from 1.27 to 0.50 , for age 2 from 0.53 to 0.19 , for age 3 from 0.24 to 0.11 . Due to the model structure predation mortality on age 4 is equal to that of age 3. Catchability on the IBTS survey is estimated to $4.0 \%$ in the first period and $4.9 \%$ in the second. Stock biomass of Norway pout was estimated to 1.36 million t in the first period and to 1.71 million $t$ in the second period.

The Coefficient of Variation (C.V.) of the observations was $20 \%$ in the first period and $8 \%$ in the second. As the observations used in the models are the mean over 5 year data points this compares roughly to CVs of $45 \%$ and $18 \%$, respectively, of individual observations. These C.V.s are model dependent and therefore likely to be under-estimates of the real uncertainty. The confidence intervals of the estimated parameters for the two periods overlap for all parameters except for fishing mortality (F).

Table 5. Maximum likelihood estimation based on IBTS cpue indices, commercial catch in numbers, and numbers eaten. Bold type represents parameters estimated. Intervals given in brackets are $95 \%$-confidence limits obtained by bootstrapping.


Statistic: C.V. of individual years observations $18 \%$

Table 6. Model sensitivity analysis. (a) Sensitivity to errors in estimated numbers consumed. (b) Sensitivity to assumed M1 of age 1 .

|  | Data basis | $\begin{gathered} \text { M1 } \\ \text { age } 2 \end{gathered}$ | $\begin{gathered} \text { M1 } \\ \text { age } 3 \end{gathered}$ | $\begin{gathered} \text { M1 } \\ \text { age } 4 \end{gathered}$ | Stock biomass |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (a) |  |  |  |  |  |
| Predation under-estimated by a factor of 2 | Model using year classes 1977-1981 | 1.85 | 2.75 | 3.21 | 2.56 mill. t |
| Predation over-estimated by a factor of 2 |  | 1.47 | 2.17 | 2.68 | 0.73 mill. t |
| Predation under-estimated by a factor of 2 | Model using year classes 1987-1991 | 2.06 | 3.07 | 4.22 | 3.07 mill. t |
| Predation over-estimated by a factor of 2 |  | 2.00 | 3.00 | 4.14 | 1.00 mill. t |

M1 of age 1

| (b) |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 0.05 | Model using year classes $1977-1981$ | 1.72 | 2.56 | 3.04 | 1.31 mill. $t$ |
| 0.2 |  | 1.79 | 2.61 | 3.09 | 1.45 mill. $t$ |
| 0.05 | Model using year classes $1987-1991$ | 2.02 | 3.03 | 4.17 | 1.59 mill. $t$ |
| 0.2 |  | 2.09 | 3.08 | 4.23 | 2.08 mill. $t$ |

Sensitivity analyses were performed to test the dependence of the estimated M1 values to uncertainties in numbers consumed. Due to e.g. large uncertainties in consumption rates used in the MSVPA for the North Sea, the numbers consumed even when averaged over 5 year classes, are expected to be estimated with large
uncertainties. Model runs with numbers consumed halved and doubled were made for each of the two time periods. In the first time period M1 on age 2 decreased from 1.74 (Table 5) to 1.47 (Table 6a) when the number consumed was halved, and increased to 1.85 when the number was doubled. For age 3 the values were 2.58,
2.17, and 2.75, respectively, and for age $43.05,2.68$, and 3.21 , respectively. In the second period the changes were much smaller. In general, the estimates of M1 were moderately sensitive to uncertainties in numbers consumed, but all values were still very high, between 1.47 and 4.22. The biomass estimates of the model runs are included in Table 6a to show that some parameters estimated or calculated were indeed very sensitive to the uncertainty in numbers consumed. The biomass estimates varied by a factor of 3 in both periods, and gave unrealistic high values of around 3 million $t$ when the doubling of numbers consumed were used. Fishing mortality and predation mortality varied to the same extent as these are closely linked to the estimated biomass.

The model's sensitivity to the assumption of M1 $=0.10$ for age 1 was also tested. Model runs with this value halved and doubled were made for each of the two time periods. The estimated M1 values for age 2 and older were insensitive to this perturbation (Table 6b). The biomass estimates were moderately sensitive and gave values between 1.31 and 2.08 million t .

The down weighing of age 4 in the commercial catches and age 5 in the IBTS data by a factor of 10 was tested on the 1987-1991 data set by varying this factor between 5 and 20. This gave less than $2 \%$ variations in the parameter estimates.

## Discussion

All the above time series and model runs give a consistent picture of a very high total mortality and one, which is increasing by age. Whether this can be an artefact is discussed below. Two possibilities could be:
(1) That catchability decreases with age for all three data series. Catchability in the IBTS survey could decrease with age if Norway pout becomes more pelagic with age. However, according to Albert (1994) Norway pout distributed close to the bottom (at most $1.2-1.5 \mathrm{~m}$ above) during day and at night up to 10 m above the bottom. Albert (1994) based his judgement on echo sounder traces. This also fits well with Raitt (1968) who stated, that available evidence suggests that Norway pout normally inhabit a zone just clear of the sea-bed. Raitt (1968) based his statement partly on stomach content data, which showed that Norway pout consumed food from that part of the water column. As most hauls at the IBTS are made during day and the net opening is around 4 m most part of the stock should be caught by the trawl, although the very low catchability estimated to be $4-5 \%$ (Table 5) indicates that avoidance is strong and this could be age dependent. However, even if Norway pout become more pelagic with age or better to avoid the trawl gear, the predators should have no problems in finding them and probably likewise the commercial fishermen.
(2) That old Norway pout migrate out of the area. The only possibility seems to be that Norway pout migrate out to the north and northwest because, according to Albert (1994) they are not found in the Norwegian Deep. Some spawning is known to take place in the North Sea, but whether this is the main spawning area is not well known. However, large numbers of old Norway pout have never been found anywhere outside the North Sea and, at least in the Faeroe Islands waters, very few Norway pout of age 3 and older are found (Jakup Reinert, Fisheries Laboratory of the Faroes, personal comm.). In light of the high mobility of the commercial fishing fleet and its ability to find commercial interesting aggregations of fish when such exists, and the large biomass of old Norway pout that would have to move to account for the estimated increase in Z , it seems unlikely that old Norway pout have migrated out of the area to any significant extent. Thus, it can be concluded that the observed increase in Z from age 1 onwards most probably is real.

As Z is the sum of F , M1 and M2 the question becomes which of these parameters is increasing with age.

Is F increasing? Fishing mortality on the age 1 and older is probably quite independent of age, as all age groups are distributed in almost the same area and they are only slightly different in size. The mesh size used in the commercial fishery for Norway pout is typically 16 mm stretched mesh and that will not allow many Norway pout of age 1 or older to escape. It can, however, not a priori be ruled out as a possibility that old Norway pout might be better to avoid being caught by being better swimmers, more "experienced", etc. However, an increase in Z as a function of F for older fish would be an explanation if the converse were true. Furthermore, if F is the only reason for the increase in $Z$, $F$ must be higher than the difference between $Z$ of age 2 and $Z$ of age 1, i.e. about 1 for age 2 and even higher for age 3. Such high F values are unrealistic. Lastly, if F really increases with age, then the older ages in the commercial catches should be over-represented compared to the younger age groups and the real Z must then be even much higher than estimated from the commercial data. The Z values calculated from the commercial catch data should then be smaller than Z values calculated from the survey data, which is not the case (Table 4). Thus, it can be ruled out that increase in F by age is the main explanation for increases in Z by age. Thus natural mortality (M) must increase with age.

Predation mortality (M2) generally decreases with age (e.g. Sparholt 1994) and the high Z values from the time series of numbers consumed supports this feature, because it indicates that old fish are under-represented compared to the other two data series. Because $\mathrm{M}=\mathrm{M} 1+\mathrm{M} 2$ it can be concluded that it must be the
residual or non-predation natural mortality (M1) which increases with age.

The result of the model runs support this conclusion. Both model runs estimate non-predation natural mortality to increase significantly with age.

Predation mortality of age 1 decreases from 1.27 for the year classes 1977-1981 to 0.50 for the year classes 1987-1991. This took place with the decrease in biomass of the three main predators of 1 group Norway pout, mature cod, whiting and saithe, from spawning stock biomass of 925000 t in 1981 to 451000 t in 1991 (ICES, 2001).

Fishing mortality is estimated to be 0.29 in the first period, but only 0.10 in the second period. The mean annual landings in weight were 351000 t in 1977-1985 where the year classes 1977-1981 were fished and 187000 t in 1987-1995 where the year classes 1987-1991 were fished. The drop in F seems consistent with the commercial landings in weight. Furthermore, according to a Danish effort data series form 1982-1998 (ICES, 1995, 2000), effort decreased from 9276 (arbitrary scale) in 1982-1985 to 3561 in 1987-1995. If 1982-1985 is representative for the entire period 1977-1985, there is a fairly good agreement between effort data and the estimated F .

The biomass estimates of 1.36 and 1.71 million t of Norway pout at 1 January, seem to be of the right order of magnitude as the total amounts consumed were 1.9 and 1.8 million $t$ in 1981 and 1991 (ICES, 1997). The commercial catch was 239000 t in 1981 and 193000 t in 1991. It might seem peculiar that the sum of the catch and the amount consumed is higher than the stock biomass, but this can be attributed to individual fish growth during the year and to recruitment.

The catchability of Norway pout on the IBTS survey is estimated to be between $4.0 \%$ and $4.9 \%$ of the fish present in the swept area. Sparholt (1990), found similarly a low catchability of Norway pout.

The estimated C.V. of the IBTS data, the commercial catches and the numbers consumed was $20 \%$ for the first period. This seems as a high uncertainty for the commercial catch data, but might be reasonable for the IBTS data (Cotter, 2001). For the numbers consumed it is difficult to say but as argued above the uncertainty might be larger than $20 \%$. It was specifically tested whether an error of a factor of 2 in both directions of the numbers consumed would give very different estimates of M1 for ages 2 and older. This was not the case as the deviations in M1 were between $1 \%$ and $16 \%$ for all ages.
It can be concluded that the non-predation or residual natural mortality, M1, of Norway pout increases with age and gets very high on age 2 and older fish.

It can further be concluded that it has been possible to estimate M1 in a way that is consistent with all the information and knowledge currently available on Norway pout.


Figure 1. Mean IBTS CPUE (in numbers per hour) per length class and quarter of the year in 1991-1996.

The slightly higher values of M1 of age 2 and older estimated for the second period than for the first period is difficult to explain and is not significant. For the time being mean M1 values for the two periods considered here are recommended for use in the MSVPA and other models, where M1 values are needed.

Norway pout length distribution data for each of the 4 quarters of the year are available for 1991-1996 from the IBTS survey (Figure 1), and these data indicate that the mortality mainly takes place between the 1st and the 2 nd quarter of the year, i.e. from before to after spawning. This points at spawning as the main reason for this mortality. Spawning mortality is well known for capelin (Mallotus villosus), another North Atlantic fish species. According to Burton and Flynn (1998) and Huse (1998) almost all male capelin dies after spawning. Tanasichuk (2000) finds that surplus energy requirements for gonad recrudescence appear to cause the death of adult Pacific herring (Clupea pallasi). To the knowledge of the present authors there are no direct records of evidence of
multiple spawning in individual Norway pout. It would be interesting to investigate whether histological analyses of the gonads of North Sea Norway pout can support or reject the spawning mortality hypothesis. For capelin, Bakke and Bjørke (1973) report dead fish in demersal trawling on the spawning grounds, but this has never been recorded for North Sea Norway pout. Demersal trawling for dead North Sea Norway pout at the spawning grounds could also be considered in order to evaluate the hypothesis.

## Acknowledgements

We thank ICES for permission to use data contained in ICES fish stock assessment reports and in the IBTS database. We also thank Ron Tanasichuk as referee and one anonymous referee for their valuable help with the manuscript and Morten Vinther for extracting estimates of number of Norway pout consumed by the predators included in the MSVPA model.

## References

Albert, O. T. 1994. Biology and ecology of Norway pout (Trisopterus esmarki Nilsson, 1855) in the Norwegian Deep. ICES Journal of Marine Science, 51: 45-61.
Anon. 1998. Feeding ecology of the North Sea Fish with emphasis on the data base of the "Stomach Sampling Project 1991" for use in multispecies assessment. EU Final Report Contract AIR3-CT94-2410, March 1995-February 1998.

Anon. 1999. Final Report EU contract no. 95/69: "Input of Historic IBTS Data". DLO-Netherlands Institute for Fisheries Research. RIVI-DLO report c014/99.
Bailey, R. S., and Kunzlik, P. A. 1984. Variation in growth and mortality rates of Norway pout Trisopterus esmarkii (NILSSON). ICES CM 1984/G: 70.
Bakke, S., and Bjørke, H. 1973. Diving observations on Barents Sea capeling at the spawning grounds off northern Norway. Fiskeridirektoratets Skrifter Serie Havundersøkelser, 16: 140-147.

Baranov, F. I. 1918. On the question of the biological basis of fisheries. Nauchnyi Issledovatelskii Ikhtiologicheskii Institut Isvestia, 1: 81-128 (In Russian).
Burton, M. P. M., and Flynn, S. H. 1998. Differential postspawning mortality among male and female capelin (Mallotus villosus Müller) in captivity. Canadian Journal of Zoology, 76: 588-592.
Cotter, A. J. R. 2001. Intercalibration of North Sea International Bottom Trawl Surveys by fitting year-class curves. ICES Journal of Marine Science, 58: 622-632.
Haddon, M. 2001. Modelling and quantitative methods in fisheries. Chapman and Hall/CRC. 406 pp.
Huse, G. 1998. Sex-specific life history strategies in capelin (Mallotus villosus)? Canadian Journal of Fisheries and Aquatic Science, 55: 631-638.
ICES 1995. Report of the Working Group on the Assessment of Norway pout and Sandeel. ICES CM 1995/Assess: 57.
ICES 1997. Report of the Multispecies Assessment Working Group. ICES CM 1997/Assess: 16.
ICES 1999. Report of the International Bottom Trawl Survey Working Group. ICES CM 1999/D: 2.
ICES 2000. Report of the Working Group on the Assessment of Demersal Stocks in the North Sea and Skagerrak. ICES CM 2000/ACFM: 7.
ICES 2001. Report of the ICES Advisory Committee on Fisheries Management. ICES Cooperative Research Report. No 242. ISSN 1017-6195.
Poulsen, E. M. 1968. Norway pout: Stock movements in the Skagerrak and the north-eastern North Sea. Rapports et Proces-Verbaux des Reunions du Conseil International pour l'Exploration de la Mer, 158: 80-85.
Raitt, D. F. S. 1968. The population dynamics of Norway Pout in the North Sea. Marine Research, No. 5: 1-23.
Shepherd, J. G. 1999. Extended survivors analysis: An improved method for the analysis of catch-at-age data and abundance indices. ICES Journal of Marine Science, 56: 584-591.
Sparholt, H. 1990. An Estimate of the total biomass of fish in the North Sea. Journal du Conceil International pour l'Exploration de la Mer, 46: 200-210.
Sparholt, H. 1994. Fish Species Interactions in the Baltic. Dana, 10: 131-162.
Sparre, P. 1991. Introduction to multispecies virtual population analysis. ICES Marine Science Symposia, 193: 12-21.
Ursin, E. 1963. On the Seasonal Variation of Growth Rate and Growth Parameters in Norway Pout (Gadus esmarki) in the Skagerrak. Med. Meddelser fra Danmarks Fiskeri- og Havundersøgelser, 4: 17-29.

