## SHORT SLEEP DURATION AND PSYCHOLOGICAL DISTRESS IN YOUNG ADULTS

# Short Sleep Duration in Prevalent and Persistent Psychological Distress in Young Adults: The DRIVE Study 

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#### Abstract

Objectives: Young people are sleeping less. Short sleep duration has a range of negative consequences including a hypothesized link with psychological distress, which has yet to be studied Design: Prospective cohort study Setting: Community-based sample from Australia Participants: Twenty thousand $(20,822)$ young adults (aged 17-24) identified through the state vehicle licensing authority. A random sample ( $\mathrm{n}=$ 5000) was approached for follow-up 12-18 months later, with 2937 providing full data.

Main Outcome Measure: Psychological distress, determined by a Kessler 10 score > 21, at baseline; and as both onset and persistence of distress at follow-up. Results: Shorter sleep duration was linearly associated with prevalent psychological distress: relative risk (RR) 1.14 ( $95 \% \mathrm{Cl} 1.12$ to 1.15). Only the very short (<5h) sleepers among those not distressed at baseline had an increased risk for onset of psychological distress (RR 3.25 [ $95 \%$ CI 1.84, 5.75]). Of 945 cohort participants reporting psychological distress at baseline, $419(44 \%)$ were distressed at follow-up. Each hour less of sleep increased the risk of psychological distress persisting after adjustment for potential confounding variables: RR 1.05 ( $95 \% \mathrm{Cl} 1.01$ to 1.10). Long sleep duration showed no association with distress at any time point. Conclusions: Self-reported shorter sleep duration is linearly associated with prevalent and persistent psychological distress in young adults. In contrast, only the very short sleepers had a raised risk of new onset of distress. Different approaches to sleep duration measurement yield different results and should guide any interventions to improve subjective sleep duration in young adults.


Keywords: Sleep, psychological distress, cohort, young adult
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## ABOUT 25\% OF YOUNG ADULTS ARE AFFECTED BY PSYCHOLOGICAL DISTRESS, A MEASURE OF POOR MENTAL HEALTH COMPRISING COMMON SYMPTOMS

 of low mood and anxiety, at any one time. ${ }^{1}$ A considerable proportion of those with distress go on to develop more severe mental disorders, recurrent episodes ${ }^{2}$ and the associated negative sequelae, including suicide attempts and welfare dependence. ${ }^{3,4}$ There appears to have been an increase in the rates of psychological distress in young adults over the past few decades which remains unexplained. ${ }^{5}$ However the persistence and recurrence of such distress in young adults is a strong predictor of a more chronic course of depression ${ }^{6}$ and as such a target for early intervention approaches, as the majority of these problems emerge in the years following puberty. ${ }^{7}$There are a great range of social, psychological, and biological shifts occurring in modifiable risk factors for mental disorder over the adolescent years, including the onset of substance use and changes in physical activity. ${ }^{8}$ One potential risk factor that has attracted recent attention is sleep disturbance, of which short sleep duration is a significant component. Sleep duration

[^0]decreases from birth, ${ }^{9}$ such that by age 17 years, it averages 7.68 h sleep on weeknights, a figure that falls by half an hour by age 24 . On average, adolescents and young adults sleep longer on weekends, but many young people frequently get less sleep than needed over the week. ${ }^{10}$ Furthermore there is evidence of temporal trends towards shorter sleep duration: In Australia young people report getting a half-hour less sleep per night over the past decade or so, particularly in those from lower socioeconomic groups. ${ }^{11}$ In the US, an increase in the proportion of people reporting short sleep is reported by all groups, but again particularly in those from disadvantaged social groups. ${ }^{12}$ Shorter sleep duration in younger age groups is associated with a number of factors, including increased television viewing, computer gaming, and internet use, ${ }^{13}$ all of which appear to be becoming more commonplace and associated with later bedtimes while waking times remain stable. ${ }^{13}$

Subjective short sleep duration is a common clinical complaint amongst people with mental disorder and a component of several diagnoses. Self-reported short sleep has been associated with psychological distress in cross-sectional studies in general adult populations, ${ }^{12,14}$ although the association of polysomnographic measured sleep duration with distress is weaker. ${ }^{14}$ Short sleep duration and later bedtimes have been shown very recently to be associated with current depressed mood and suicidal ideation in adolescents. ${ }^{15}$ Sleep disturbance in adults, often characterized as insomnia, has consistently been shown to increase the likelihood of the onset of future mental disorder in those currently without disorder ${ }^{16-18}$ and confirmed recently in both
self-report and polysomnographic data in the Wisconsin Sleep Cohort. ${ }^{19}$ None of these longitudinal studies of sleep disturbance and future mental disorder included participants aged under 20, yet this is an age group in which major changes take place both in rates of mental disorder and sleep patterns. ${ }^{7}$ All of the above studies compared a group of very short sleepers with the rest of the sample and so could not evaluate to what extent any observed association held across the range of sleep duration. This has implication for targeting interventions. Further there are no prospective studies evaluating the impact of short sleep duration in younger adults on the onset or persistence of the broader phenotype of psychological distress. Nor are there any studies addressing the impact of long sleep duration, although in clinical samples of people with psychiatric disorder hypersomnia is not uncommon. This study aimed to answer whether self-reported sleep duration, defined both categorically and continuously, is independently associated with psychological distress and predicts its onset and persistence in young adults.

## METHODS

A prospective cohort study examining risk factors for young driver crash and injury recruited 20,822 young Australians, aged 17-24 years; full baseline data on sleep duration and psychological distress were available for $94.4 \%(19,648)$. The methods and characteristic of the sample are described in detail in an earlier publication. ${ }^{20}$ In brief, drivers aged 17-24 years holding their first-stage provisional motor-vehicle driver's license in the State of New South Wales (NSW), Australia, were invited to complete a confidential survey that could be completed either on-line or via a printed questionnaire. The survey included questions on psychological distress and sleep (described below), as well as a range of other demographic and driving related risk factors, such as sensation seeking, risky driving behaviors, and substance use.

## Measures

Subjective sleep duration was assessed through self-report of the number of hours slept on weekday and weekend nights separately over the past month. The weighted average of weekday and weekend sleep hours was used to determine average sleep duration. Such a measure is a standard assessment of perceived hours slept used in a number of large epidemiological studies that have reported the consequences of sleep duration, such as the National Health and Nutrition Examination Survey, ${ }^{21}$ and shown there to be associated with objectively reported bedtimes in this age group. ${ }^{15}$ For the descriptive univariate analyses, this was then grouped into an ordinal variable. Only 157 participants reported sleeping $>11 \mathrm{~h}$ on average; and as the same proportion ( $31 \%$ ) of this group had psychological distress as the $10-$ to $11-\mathrm{h}$ group, these groups were merged. Participants reporting impossible sleep duration, e.g., 150 hours per weekend $(\mathrm{n}=297)$ or missing sleep duration data $(\mathrm{n}=795)$ were excluded. The a priori main analyses were based upon log binomial models using a continuous measure of average sleep duration to retain information.

Several potential confounders were evaluated. Drug use was measured using 2 questions from the Centers for Disease Control and Prevention (CDC) Youth Risk Behavior Surveillance System on the use of marijuana (cannabis) and other recreation-
al drugs in the past 12 months (coded as ever /never). ${ }^{22}$ Alcohol use was measured using a 3 -item subscale of the Alcohol Use Disorders Identification Test (AUDIT-C), with a score $>6$ indicating harmful use. ${ }^{23}$

Self-harm was assessed with an open-ended question as in previous Australian adolescent and young adult research: "In the past 12 months have you deliberately hurt yourself or done anything that you knew might have harmed you or even killed you? (yes/ no), followed by a clarification question "If yes, what was it that you did?" These questions were based on the Beck Suicide Inventory and have been previously used in studies in young adults in Australia. ${ }^{24,25}$ Two trained research assistants then coded the open responses to evaluate whether the reported action constituted "deliberate self-harm" (DSH), rather than engagement in such commonly reported risk-taking activities as driving over the speed limit. Initial training and review of the first 100 positive responses were conducted by a consultant psychiatrist. All discrepancies were reviewed in further training before completing the rest of the sample. The inter-rater reliability of the trainers for a post-training random sample of $20 \%$ of these codes was a $\kappa$ of 0.97 .

Demographic and other routine data were also recorded and categorized as gender, age ( $17,18,19+$ ), and occupation (secondary /tertiary student, working, and not working). Given the purported effects of local environment and socioeconomic status upon mental disorders, this included a measure of socioeconomic status derived from Socio-Economic Indexes for Areas (SEIFA) scores based on small-area statistics from the Australian 2001 population census.

## Outcome

Psychological distress was assessed using the Kessler 10 (K10). ${ }^{26}$ This 10 -item self-report measure asks: "The following questions are about your feelings in the past 4 weeks: in the past 4 weeks how often did you feel tired out for no reason?" Further questions ask about feeling: nervous, hopeless, restless, depressed, that everything is an effort, very sad, and worthless, but no somatic features. There are 5 response options ranging from "None of the time" to "All of the time," and scores range from 10 to 50 . Those scoring $>21$ were considered to be experiencing significant levels of psychological distress and likely to represent cases of mental disorder in Australian populations, as reported in government surveys of young people. ${ }^{1}$

## Resurvey

A random sample of 5000 of the initial respondents was recontacted 12-18 months after completing the baseline survey by text message, email, and/ or letter. It was not possible to resurvey the full cohort due to resource limitations. Resurvey respondents were invited to complete a second questionnaire and were reimbursed for their time with a cinema ticket. A total of 3344 (67\%) resurveyed individuals responded, of whom 2937 had full baseline data and follow-up K10 data. Of these, 1992 (67.8\%) did not report psychological distress at baseline and responded with one of 2 outcomes: a Kessler-10 (K10) score $>21$ at resurvey, termed "onset of psychological distress," or repeated lower levels of distress $(\mathrm{K} 10 \leq 21)$. Of the 2937 respondents with complete data, 945 ( $32.2 \%$ ) had high levels of distress at baseline and were coded into one of 2 outcomes: again reporting a K10 score $>21$ at resurvey, termed "persistent psychological
distress," or of "remission" to lower levels of distress (K10 $\leq$ 21). Those not responding to the resurvey were more likely to be male ( $43 \%$ vs $34 \%$ in follow-up sample), have finished education ( $17 \%$ vs $11 \%$ ), and be substance misusers (e.g., harmful alcohol users $20 \%$ vs $13 \%$ ) and high sensation seekers ( $43 \%$ vs $35 \%$ ) ( P values for each $\chi^{2}$ test $<0.001$ ). Mean sleep duration did not differ between the responders (7.7 h) and non-responders ( 7.6 h ). There was no response bias from psychological distress, self-harm reporting, socioeconomic status, or age.

The University of Sydney Human Research Ethics Committee and the NSW Health Ethics Committee approved the study.

## Analysis

The associations of demographic, behavioral and psychological characteristics with categories of sleep duration and psychological distress were examined using $\chi^{2}$ tests, and for trend if necessary, in the baseline data. The association of sleep duration with prevalent, onset and persistence of psychological distress was analyzed using both the grouped sleep duration for interpretation and using the continuous measure of average sleep duration. The modal sleep duration of $\geq 8$ to 9 hours was used as the reference group for descriptive analyses.

After confirmation of the log-linearity of the relationship, log binomial models examining the association of sleep duration (using an ordinal measure of hours) with psychological distress were constructed, first unadjusted and then adjusting sequentially for a priori demographic (model 1) and behavioral (model 2) factors. Given the log linear effect of the sleep values used for the ordinal variable then this should be an appropriate summary of the log linear effect of categorical groups.

This analysis was repeated for the association of sleep duration with onset of psychological distress in the sample of those with no distress at baseline $(\mathrm{n}=1992)$ and persistence in those with baseline psychological distress at baseline $(\mathrm{n}=945)$, who formed 2 mutually exclusive cohorts. Baseline severity (K10 score) was additionally entered (as quartiles) to allow for the consistent finding in studies of baseline symptom severity as a predictor of poorer outcome. A third model weighted those variables associated with non-response, with weights based upon the degree of potential bias. As an additional check, the final model was also adjusted for an offset according to length of time in the study.

## RESULTS

## Association of Sleep Duration with Current Psychological Distress

Overall, $18 \%$ of the 19,648 with full data in this study reported sleeping $<7 \mathrm{~h}$ per night on average, and a further $30 \%$ slept between 7 h and 8 h . On average this sample slept 20 min more on weekends ( 8.1 h , SD 1.98) than during the week (7.8 h, SD 1.35). Almost one-third ( $32.5 \%$ ) scored $>21$ on the K10, indicating high levels of current psychological distress.

Shorter sleep duration was associated with psychological distress ( $\chi^{2}$ test for trend $\mathrm{P} \leq 0.0001$ ) (Table 1). Over half of those reporting $<6 \mathrm{~h}$ sleep per night had high levels of current psychological distress, compared to about one quarter of those sleeping the recommended level in young adults of $8-9 \mathrm{~h}$ per

Table 1-Association of self-reported average sleep duration with psychological distress in the baseline sample of 19,648 young adults

| Average sleep duration (categories) | High or very high psychological distress K10 > 21 |  | No or moderate psychological distress K10 $\leq 21$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} \mathrm{N} \\ (6386) \end{gathered}$ | $\begin{gathered} \% \\ \text { (row) } \end{gathered}$ | $\begin{gathered} \mathrm{N} \\ (13,262) \end{gathered}$ | $\begin{gathered} \% \\ \text { (row) } \end{gathered}$ | Relative risk (RR and 95\%CI) |
| $\geq 10 \mathrm{~h}$ | 317 | 30.6 | 718 | 69.4 | 1.16 (1.05, 1.28) |
| $\geq 9$ to <10 h | 676 | 25.3 | 2001 | 74.7 | 0.96 (0.89, 1.03) |
| $\geq 8$ to <9 h | 1653 | 26.4 | 4608 | 73.6 | REF |
| $\geq 7$ to <8 h | 2045 | 34.1 | 3955 | 65.9 | 1.29 (1.22, 1.36) |
| $\geq 6$ to $<7 \mathrm{~h}$ | 1154 | 43.2 | 1520 | 56.8 | 1.63 (1.54, 1.74) |
| $\geq 5$ to < 6 h | 329 | 51.5 | 310 | 48.5 | 1.95 (1.79, 2.13) |
| $\leq 5 \mathrm{~h}$ | 212 | 58.6 | 150 | 41.4 | 2.22 (2.02, 2.44) |
| Average sleep | ration (d | easing | urs) |  | 1.14 (1.12, 1.15) |

night: (relative risk [RR] for $\leq 5 \mathrm{~h}, 2.22$ [ $95 \% \mathrm{CI}$; 2.02, 2.44], > 5 to $<6$ h [RR 1.95], $95 \% \mathrm{CI} ; 1.79,2.13$ ). Modeling the continuous measure of sleep duration rather than using grouped data demonstrated a linear association of reported average sleep duration with psychological distress (RR 1.14; 95\%CI 1.12,1.15), indicating $14 \%$ greater risk of higher levels of distress for every hour less slept. There was no such association with longer sleep duration.

Self-reported sleep duration was associated with a number of baseline characteristics (Table 2): Older age, being employed, recent deliberate self-harm, and reported use of marijuana, other drugs, and drinking at harmful levels were all associated with shorter sleep duration ( $\mathrm{P}<0.001$ in each case for test for trend). Male gender and being unemployed showed U-shaped associations with sleep duration. In this sample, high psychological distress was associated with female gender ( $40 \%$ vs $28 \%$ in males), unemployment ( $33 \%$ vs $28 \%$ in the employed), drug use ( $45 \%$ vs. $32 \%$ ), harmful alcohol use ( $38 \%$ vs $32 \%$ ), high sensation-seeking behavior ( $44 \%$ vs $22 \%$ in lowest category) and recent deliberate self-harm ( $70 \%$ vs $31 \%$ ) ( $\mathrm{P}<0.001$ in each case). The area level socioeconomic status score had no such association ( $\mathrm{P}=0.8779$ ).

## Effect of Short Sleep Duration upon Onset of Psychological Distress

Amongst the 1992 participants in the resurvey who did not report high levels of stress at baseline and had follow up data, 239 (12\%) reported high levels of distress one year later (Table 3). Although the ordinal categorization of sleep duration suggested a reverse J-shaped curve, with only those in the very shortest category having a higher risk (RR 3.25 [(95\% CI 1.84, 5.75]), when modeling the data, there was a linear association (RR 1.12 [ $95 \%$ CI $1.01,1.24]$ ) of shorter sleep duration increasing the likelihood of onset of psychological distress. There was no effect of long sleep duration upon new onset of distress.

In a multivariable model (Table 4), although adjusting for demographic factors had no effect, the linear association was apparently confounded by subthreshold levels of distress and behavioral factors ( 1.06 [ $95 \%$ CI $0.96,1.17]$ ). Of note was that only these lower levels of distress symptoms (and possibly, being unemployed) predicted the onset of distress in these models.

## Effect of Sleep Duration upon Persisting Psychological Distress

Of the 945 participants in the resurvey who reported a K10 score $>21$ at baseline, 419 ( $44 \%$ ) of these reported experiencing high or very high levels psychological distress one year later. These were considered the "persistent" cases and were
compared to those $(\mathrm{n}=526)$ whose distress had "remitted." Shorter sleep duration showed a significant effect of being associated with a greater risk of persisting psychological distress in all sleep duration groups less than the recommended reference group duration (Table 3).

Table 2-Association of demographic and psychological characteristics with sleep duration at baseline in 19,735 young adults

| Characteristic | Sleep hours |  |  |  |  |  |  |  |  |  | Total |  | $\begin{gathered} \mathrm{P} \\ \text { value } \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 0-5 |  | $>5$ to $<7$ |  | $\geq 7$ to < 8 |  | $\geq 8$ to < 9 |  | $\geq 9$ |  |  |  |  |
|  | $\mathrm{n}=363$ | \% (col) | $n=3328$ | \% (col) | $n=6017$ | \% (col) | $n=6291$ | \% (col) | $\mathrm{n}=3736$ | \% (col) | $n=19,735$ | \% (col) |  |
| 17 | 132 | 36.36 | 1378 | 41.41 | 2895 | 48.11 | 3317 | 52.73 | 1916 | 51.28 | 9638 | 48.84 |  |
| 18-19 | 154 | 42.42 | 1379 | 41.44 | 2192 | 36.43 | 2206 | 35.07 | 1390 | 37.21 | 7321 | 37.10 |  |
| $\geq 20$ | 77 | 21.21 | 571 | 17.16 | 930 | 15.46 | 768 | 12.21 | 430 | 11.51 | 2776 | 14.07 | < 0.0001 |
| Sex |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Female | 184 | 50.69 | 1890 | 56.79 | 3339 | 55.49 | 3364 | 53.47 | 1977 | 52.92 | 10754 | 54.49 |  |
| Male | 179 | 49.31 | 1438 | 43.21 | 2678 | 44.51 | 2927 | 46.53 | 1759 | 47.08 | 8981 | 45.51 | 0.0012 |
| Occupation |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 1. High school student | 123 | 34.07 | 1360 | 41.40 | 2992 | 50.34 | 3398 | 54.87 | 1934 | 52.67 | 9807 | 50.41 |  |
| 2. TAFE/Uni student | 104 | 28.81 | 1049 | 31.93 | 1673 | 28.15 | 1557 | 25.14 | 949 | 25.84 | 5332 | 27.41 |  |
| 3. Employed | 95 | 26.32 | 729 | 22.19 | 1092 | 18.37 | 1014 | 16.37 | 544 | 14.81 | 3474 | 17.86 |  |
| 4. No work / benefits | 39 | 10.80 | 147 | 4.47 | 186 | 3.13 | 224 | 3.62 | 245 | 6.67 | 841 | 4.32 | < 0.0001 |
| Alcohol harmful drink |  |  |  |  |  |  |  |  |  |  |  |  |  |
| No | 271 | 74.66 | 2731 | 82.06 | 5250 | 87.27 | 5558 | 88.39 | 3315 | 88.76 | 17125 | 86.80 |  |
| Yes | 92 | 25.34 | 597 | 17.94 | 766 | 12.73 | 730 | 11.61 | 420 | 11.24 | 2605 | 13.20 | < 0.0001 |
| Marijuana use |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Never | 283 | 78.39 | 2672 | 80.48 | 5184 | 86.30 | 5499 | 87.56 | 3311 | 88.89 | 16949 | 86.07 |  |
| Ever | 78 | 21.61 | 648 | 19.52 | 823 | 13.70 | 781 | 12.44 | 414 | 11.11 | 2744 | 13.93 | < 0.0001 |
| Drugs use |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Never | 298 | 82.09 | 2972 | 89.55 | 5586 | 93.27 | 5941 | 94.81 | 3517 | 94.59 | 18314 | 93.18 |  |
| Ever | 65 | 17.91 | 347 | 10.45 | 403 | 6.73 | 325 | 5.19 | 201 | 5.41 | 1341 | 6.82 | < 0.0001 |
| Self-harm in past year |  |  |  |  |  |  |  |  |  |  |  |  |  |
| No | 312 | 85.95 | 3103 | 93.24 | 5777 | 96.01 | 6093 | 96.85 | 3607 | 96.55 | 18892 | 95.73 |  |
| Yes | 51 | 14.05 | 225 | 6.76 | 240 | 3.99 | 198 | 3.15 | 129 | 3.45 | 843 | 4.27 | < 0.0001 |

Average
sleep
duration
(categories)
$\geq 10 \mathrm{~h}$
$\geq 9$ to $<10 \mathrm{~h}$
$\geq 8$ to $<9 \mathrm{~h}$
$\geq 7$ to $<8 \mathrm{~h}$
$\geq 6$ to $<7 \mathrm{~h}$
$>5$ to $<6 \mathrm{~h}$
$\leq 5 \mathrm{~h}$

Average sleep duration (decreasing hours)
"Onset of psychological distress" defined as a K10 score > 21 at follow-up but < 21 at baseline; "Persistent psychological distress" defined as a K10 score > 21 at both baseline and follow-up

Again there was a linear association of shorter sleep duration being associated with persistent psychological distress (RR for each hour less 1.08 [95\%CI 1.04, 1.13]). The $\log$ binomial model results for persistence are presented in Table 5: the association was only minimally attenuated by adjusting for a range of potential confounders (RR 1.05 [95\%CI $1.01,1.10]$ ). In the fully adjusted models, being older, having higher initial psychological distress severity, and reporting deliberate selfharm in the year prior to the baseline assessment also increased this risk. Behavioral factors such as harmful levels of drinking and marijuana use showed no independent effect on the persistence of psychological distress.

Weighting those variables associated with potential attrition bias (gender, higher education, sensation seeking, and substance misuse) by the degree of non-response in the resurvey, individually and in combination, had no substantial effect upon the adjusted association of sleep duration with either onset or persistence of psychological distress (Tables 4 and 5). Forcing other factors such as socioeconomic status into the model also had no significant effect. There was no interaction with either age or gender. Finally, there was no effect of adjusting for differential amounts of time in the study.

## DISCUSSION

The results from this large prospective cohort indicate that self-reported shorter sleep duration was associated with prevalent psychological distress and was an independent risk factor for its persistence a year later, even after adjusting for severity of the distress and a range of potential confounders typically associated with poor psychiatric outcomes. Initial symptom severity, older age, and recent self-harm attempts were other factors that predicted this persistence; none of the forms of substance misuse were predictive. These results suggest that short sleep duration is a marker for a poorer outcome of psychological distress in linear fashion, independent of baseline symptom (gender, age, risk taking, and alcohol misuse) (gender, age, risk taking, and alcohol misuse)

Table 4-Multivariable longitudinal log binomial models of the association of self-reported sleep duration, measured at baseline, with onset of psychological distress 12-18 months later, sequentially adjusted for confounders ( $n=1922$ )
$\left.\begin{array}{lccc} & \begin{array}{c}\text { Model 1 } \\ \text { (+ demographics) }\end{array} & \begin{array}{c}\text { Model 2 } \\ \text { (+ behavioral factors) } \\ \text { RR (95\%Cls) }\end{array} & \begin{array}{c}\text { Model 3 }\end{array} \\ \text { (+ weighted for non-response) }\end{array}\right)$

Model 1, Demographic factors in table all simultaneously entered; Model 2, Additional behavioral factors in table all simultaneously entered; Model 3, Variables associated with non-response weighted by degree of non-response

Table 5-Multivariable longitudinal log binomial models of the association of shorter sleep duration, measured at baseline, with persistent psychological distress 12-18 months later sequentially adjusted for confounders ( $n=945$ )

|  | Model 1 <br> (+ demographics) <br> RR (95\%Cls) | Model 2 <br> (+ behavioral factors) | RR (95\%Cls) <br> (+ weighted for non-response) |
| :--- | :---: | :---: | :---: |
| Characteristics | RR (95\%Cls) |  |  |
| Sleep duration (decreasing hours) | $1.09(1.03,1.14)$ | $1.06(1.02,1.10)$ | $1.05(1.01,1.10)$ |
| Age (17 vs $\geq$ 18) | $0.70(0.57,0.86)$ | $0.71(0.59,0.86)$ | $0.70(0.58,0.84)$ |
| Gender (female vs male) | $1.17(1.00,1.37)$ | $1.12(0.97,1.31)$ | $1.14(0.99,1.31)$ |
| Occupation (vs school student) |  |  |  |
| $\quad$ Higher education | $0.98(0.79,1.20)$ | $0.97(0.81,1.16)$ | $0.95(0.79,1.14)$ |
| $\quad$ Employed | $0.97(0.75,1.25)$ | $0.94(0.74,1.19)$ | $0.91(0.73,1.14)$ |
| $\quad$ Not working or student | $1.08(0.77,1.52)$ | $0.96(0.69,1.33)$ | $0.90(0.66,1.23)$ |
| K10 baseline distress (quartiles) |  | $1.24(1.17,1.32)$ | $1.23(1.15,1.30)$ |
| Harmful drinking (yes vs no) |  | $0.88(0.71,1.11)$ | $0.97(0.84,1.11)$ |
| Marijuana use (yes vs no) |  | $1.07(0.89,1.28)$ | $1.05(0.88,1.25)$ |
| Deliberate self-harm (yes vs no) |  | $1.22(1.05,1.41)$ | $1.29(1.12,1.49)$ |

Model 1, Demographic factors in table all simultaneously entered; Model 2, Additional behavioral factors in table all simultaneously entered; Model 3, Variables associated with non-response weighted by degree of non-response
and importantly, the way sleep duration is evaluated. Objective and subjective short sleep duration are not the same thing; the latter is likely to be a proxy for a range of psychological and sociodemographic factors (including gender, socioeconomic class, and occupational demands. ${ }^{28}$ The associations between short sleep and other outcomes such as obesity are also affected differentially by whether the duration is measured subjectively or objectively. ${ }^{29,30}$ In elucidating any underlying pathological pathways, this distinction needs to be borne in mind. Inflammatory markers associated with psychiatric disorder are also differentially associated with subjective and objective sleep duration. ${ }^{31}$ Our results also showed that different approaches in coding and categorizing self-reported sleep duration can have significant implications for results, even in such a large sample; this has implications for standardizing such metrics to enable comparison between studies.

One possible explanation for an association between sleep duration and psychological distress may be that short sleep duration is just a symptom of the psychological distress being assessed. Sleep disturbance is a key symptom in mental disorders such as depression, ${ }^{2}$ although it is also a common finding in populations without mental disorder. ${ }^{10}$ Short sleep duration may be a part of the prodrome of a mental disorder ${ }^{32}$ or a residual symptom of a prior disorder ${ }^{33}$ and is a marker for recurrence of depression in the elderly. ${ }^{34}$ The attenuation of the association of short duration with onset after adjustment for subclinical symptoms severity would support this view. Without an extensive life course approach, it is difficult for studies to elucidate whether any sleep disturbance was truly "prior," part of a "prodrome," or a residual symptom of previous distress or disorder. Defining the temporal relationships between sleep disturbance and forms of psychological distress has proved complex, with bidirectional paths observed, complicated by varying definitions used for both exposure and outcome. ${ }^{35}$ Recent research on sleep disturbance and mental disorder has shifted the traditional view of sleep disturbance invariably being part of a mental disorder to one of comorbidity, ${ }^{36}$ with greater understanding of common and differential underlying biology. ${ }^{37}$ The independent effect of short sleep duration upon persistent distress may also suggest that sleep disturbance is a marker for severity or that it reflects a comorbid condition that diminishes the chance of the distress resolving. Finally it raises the possibility that the increase in levels of distress reported by young adults over the past decade may reflect temporal changes in young people's sleep patterns.

The study has a number of limitations. First, the initial DRIVE study was designed to provide robust estimates of exposure-response relationships and is not population representative. ${ }^{20}$ However this sample reported similar levels of psychological ill health to the recent Australian National Survey of Mental Health and Wellbeing, ${ }^{1}$ in which $26 \%$ of 16 - to 24 -yearolds had an ICD-10 defined mental disorder. The average sleep duration and differences between weekend and weekdays are similar to those observed in the Australian Time Use survey. ${ }^{38}$ Second, there was potential attrition bias. We weighted the binomial models for those variables associated with non-response with no effect. Further there was no association of the primary exposure in this paper (sleep duration) with attrition. Third, in terms of the outcome of persistence, we cannot say whether a person with high K10 scores at both time points was chroni-
cally distressed or whether they had fluctuations in distress, as we could only evaluate at baseline and once again after a period of a year or so. This study did not collect information on a number of potential confounders, such as psychiatric or sleep diagnoses or treatment undertaken either during or prior to the study period. However, in Australia only $23 \%$ of young people with a mental disorder have used any form of services in the past 12 months. ${ }^{1}$

Given the cost and logistical limitations of both polysomnography and actigraphy, the likelihood is that most clinicians outside of sleep clinics will rely upon self-report to identify short sleep duration. The finding that self-reported shorter sleep duration increased the likelihood of persistence of psychological distress in young adults has important clinical and policy implications. Short sleep duration in adolescents and young adults is consistently implicated in predicting poorer educational attainment, somatic and, interpersonal problems, weight gain, and future obesity, ${ }^{29,39}$ and recently with depressed mood and suicidal ideation. ${ }^{15}$ The addition of another negative outcome (persistent psychological distress and thus likelihood of future psychiatric disorder) of short sleep duration adds weight to the argument for improving sleep in this age group, as well as identifying sleep duration as a marker for early intervention.

The differential results with respect to onset and persisting psychological distress highlights tension between intervention approaches. Firstly it suggests that if prevention of onset of distress is the aim, then attempts to improve sleep should be targeted only at those in the very short duration category. Interventions in those who report even slightly longer sleep may have risks, such as establishing the sleep ruminations seen in chronic insomnia. However, attempts to improve self-reported sleep duration in those with current distress may be beneficial for across the range of sleep duration.

Although there are recent individual studies of sleep-focused interventions improving not just sleep but other psychological parameters, ${ }^{40}$ as Roberts et al. observed, "much of the remedy for adolescent sleep deprivation lies in the broader societal context of adolescent lives, particularly as related to school, leisure, and work. ${ }^{3}{ }^{39}$ Modifying, or ameliorating the impact of, modern lifestyles that promote sleep deprivation in young people, ,, ${ }^{8,11,13}$ such as starting school later and restricting late night internet access and other bedtime prolonging activities is a challenging task. Health education in schools and colleges enable young people to be better informed about their choices. Such interventions should be guided by whether they are universal (everyone), indicated (subclinical symptoms), or selected (higher risk groups), ${ }^{41}$ as there may be untoward consequences of targeting those with and without current psychological distress in the same way.

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## DISCLOSURE STATEMENT

This was not an industry supported study. Dr. Glozier has been a member of the advisory panel of Sanofi-Aventis and has
participated in speaking engagements for CSL Laboratories. Dr. Hickie was Chief Executive Officer and Clinical Adviser of beyondblue, the Australian National Depression Initiative. He has lead projects for health professionals and the community supported by government, community agencies and drug industry partners (Wyeth, Eli Lily, Servier, Pfizer, Astra Zeneca) on identifying and managing depression and anxiety. He has served on professional advisory boards convened by the drug industry in relation to specific antidepressants, including nefazodone, duloxetine and desvenlafaxine and participated in a multicentre clinical trial supported by Servier. The other authors have indicated no financial conflicts of interest.

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