

Visit-to-Visit and Ambulatory Blood Pressure Variability as Predictors of Incident Cardiovascular Events in Patients With Hypertension

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BACKGROUND

Visit-to-visit blood pressure variability (BPV) has been shown to be a prognostic indicator in hypertensive patients. We designed this study to clarify the impacts of clinic and ambulatory BPV in predicting cardiovascular disease (CVD).

METHODS

We performed ambulatory BP monitoring (ABPM) in 457 hypertensive patients. Visit-to-visit BPV and ambulatory BPV were calculated as the SDs of clinic BP, awake BP, and sleep BP. The mean age of the subjects was 67.0 ± 9.2 years, and they were followed for 67 ± 26 months. Stroke, myocardial infarction, and sudden cardiac death were defined as *Hard CVD* events, and these plus angina, heart failure, and other CVDs were defined as *All CVD* events. Multivariable Cox hazard regression models predicting CVD events were used to estimate the adjusted hazard ratio (HR) and 95% confidence interval (CI) for different measures of BPV with adjustment for significant covariates.

RESULTS

In multivariable analyses, the BPV of clinic systolic BP (SBP) was an independent predictor for All CVD events (HR, 2.20; 95% CI, 1.25–3.88; $P < 0.01$), but not for Hard CVD events ($P = 0.20$). On the other hand, the BPV of sleep SBP was an independent predictor for Hard CVD events (HR, 2.21; 95% CI, 1.08–4.53; $P = 0.03$), but not for All CVD events ($P = 0.88$). Diastolic BPV exhibited the same pattern.

CONCLUSIONS

These findings suggest that visit-to-visit BPV and ambulatory BPV are separately useful in predicting cardiovascular outcomes.

Keywords: ambulatory blood pressure monitoring; ambulatory blood pressure variability; blood pressure; cardiovascular disease; clinic blood pressure variability; hypertension

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It has been more than a decade since blood pressure (BP) variability has been recognized as both a marker and risk factor for cardiovascular disease (CVD).^{1–5} BP variability (BPV) reflects stiffening of the blood vessels,⁶ sympathetic nerve activation,⁷ impaired baroreflex sensitivity,⁸ and other intrinsic and social factors.⁹ In recent years, along with the development of 24-h ambulatory BP monitoring (ABPM) and home BP monitoring, various types of BPV have been shown to be associated with cardiovascular risk. Variability in ambulatory BP has been shown to be associated with cardiovascular events in subjects with hypertension^{1,2,5,10} and diabetes.¹¹ Home BPV, especially day-to-day BPV, has been shown to be associated with adverse cardiovascular prognosis in hypertensive patients.¹² Diurnal BP variations, such as morning BP surge¹³ and a riser pattern,^{14–16} have been shown to be associated with future stroke events in hypertensive patients.

Short-term BPV, such as episodic hypertension, has not been regarded as important for the assessment of hypertension.¹⁷ When clinic BP has been compared with ABPM or home BP monitoring, it has been found inferior to both these measures for predicting target organ damage and cardiovascular outcomes.^{16,18–22} ABPM and home BP monitoring are useful methods to assess BPV because they can provide large quantities of BP data, but one disadvantage of these methods is that there is currently no large database of ABPM and home BP monitoring data, such as a large-scale clinical trial. In contrast, visit-to-visit BPV in large-scale clinical trials has been shown to be associated with cardiovascular risk. Rothwell²³ has demonstrated that visit-to-visit BPV may be an even stronger predictor than ambulatory BPV in hypertensive subjects. However, there is still no appropriate dataset for comparing visit-to-visit clinic BPV and ambulatory BPV as predictors of cardiovascular events. In the present study, we tested the hypothesis that clinic BPV would be superior to ambulatory BPV in predicting CVD.

METHODS

This study was performed in a sample of 457 asymptomatic subjects who were seen for the evaluation and management of

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hypertension in general internal medicine clinics at three institutes in Japan: one clinic and two hospitals that participated in the Karatsu-Nishiarita study.²⁴

Subjects and definitions. During the period of recruitment, from 1996 to 2002, hypertensive or possibly hypertensive patients who agreed to undergo ABPM were enrolled consecutively in the clinics. The mean age was 67.0 ± 9.2 years (range 33–88 years) and there were 172 men and 285 women. Hypertension was diagnosed when the clinic systolic BP (SBP) was ≥ 140 mm Hg and/or the diastolic BP was ≥ 90 mm Hg on at least two occasions according to current guidelines,²⁵ or when there was a previous diagnosis of hypertension with current antihypertensive medication use. Subjects took no antihypertensive medication for a minimum of 7 days before the ABPM, and more than 95% took no medications during the 14 days preceding the ABPM study. Type 2 diabetes was diagnosed according to the guidelines of the American Diabetes Association²⁶ or when there was a previous diagnosis with current use of antidiabetic medication. We excluded patients with type 1 or secondary diabetes, renal dysfunction (serum creatinine >1.9 mg/dl), hepatic damage (aspartate aminotransferase/alanine aminotransferase $>$ twice their upper limits), ischemic heart disease or other cardiac diseases, congestive heart failure, arrhythmias (including atrial fibrillation), stroke (including transient ischemic attacks), or other major concomitant non-CVD. Ischemic heart disease and stroke were checked by attending doctors with medical records, physical examinations, and laboratory and radiological findings, and those were reviewed by the investigators. Body mass index was calculated as weight (kg)/height² (m²). Current smoking status was defined as smoking within the past year. This study was approved by the institutional review board of each participating hospital or clinic. All the subjects studied were ambulatory and gave informed consent for the study.

Clinic BP measurement. At baseline, three clinic BP readings were taken on each of at least two visits (six readings in all) after at least 5 min of rest in the sitting position, which included before or after being fitted with an ABPM in subjects who stopped medication for ABPM. “Baseline clinic BP” was defined as the average BP from two different visits in untreated subjects. For treated subjects, the antihypertensive medications were stopped for 14 days and the BP readings during the untreated period were used. After the baseline assessment, clinic BP was measured with a mercury sphygmomanometer every month. Clinic BP was measured three times after a 5-min rest and the average of the second and third readings was recorded. All available clinic BP assessments between the baseline and the end of follow-up were entered into the database for each subject, and their average (“mean clinic BP”) was used in subsequent analyses. The number of postbaseline clinic BP assessments ranged from 1 to 78 (average \pm s.d., 36.5 ± 22.6 assessments per subject), and the five subjects who had only one assessment were excluded from the analysis.

ABPM. Noninvasive ABPM was performed on a weekday with an automatic system (either TM2421 or TM2425; A&D, Tokyo, Japan) which records BP, using the oscillometric method and pulse rate every 30 min for 24 h. These devices have been previously validated.²⁷ Awake and sleep times were defined based on written diaries of the patients recorded during ABPM. Mean awake and sleep levels of SBP and diastolic BP were computed and the nocturnal BP fall (%) was calculated as (awake SBP–sleep SBP)/awake SBP.

Follow-up and events. During the follow-up period, standard medical therapy was performed based on current guidelines.^{28,29} The subjects’ medical records were reviewed annually for the purpose of identifying incident CVD. When annual contact was not sufficient, a research assistant made phone calls for missing subjects. Attending doctors reviewed all the medical records, blinded to ABPM data, and the authors evaluated the endpoints based on the following criteria. Strokes and cardiac events were diagnosed by the physician caring for the patient at the time of the event, and independent neurologists or cardiologists reviewed the cases and confirmed the diagnosis by referrals or medical records including brain computed tomography or magnetic resonance imaging. Myocardial infarction was diagnosed based on the American Heart Association criterion of “definite” myocardial infarction.³⁰ Stroke was diagnosed on the basis of sudden onset of a neurological deficit that persisted for >24 h in the absence of any other disease process that could explain the symptoms.¹⁵ Stroke events included ischemic stroke (cerebral infarction and cerebral embolism), hemorrhagic stroke (cerebral hemorrhage and subarachnoid hemorrhage), and undefined types of stroke. Sudden cardiac death was defined as sudden unexpected death due to cardiac causes within 1 h after the onset of symptom.

A follow-up examination was performed in all participants in the Karatsu-Nishiarita study from March 2004 to October 2007. The mean follow-up period was 66 ± 27 months. We defined three outcomes as *Hard* CVD events: stroke ($n = 26$), fatal or nonfatal myocardial infarction ($n = 5$), and sudden cardiac death ($n = 3$). Participants who became dependent in their daily living ($n = 6$), those who died or suffered from noncardiovascular causes such as malignant disease, accident, or neurologic disorders ($n = 16$), and those who moved or changed their telephone number ($n = 4$) were censored as of the time such events took place ($n = 26$ subjects in total). In this study, angina confirmed by a significant stenosis by coronary angiography ($n = 7$), congestive heart failure requiring hospitalization ($n = 9$), end-stage renal disease requiring hemodialysis ($n = 2$), peripheral artery disease confirmed by objective tests such as ankle-brachial index <0.9 ($n = 3$)³¹ and transient ischemic attacks requiring hospitalization ($n = 3$), in which the neurological deficit was completely cleared within 24 h,¹⁵ were treated as *Soft* CVD events. These CVD events were combined with *Hard* CVD events to define *All* CVD events. When subjects did not visit the clinics, we interviewed them by telephone.

Statistical analyses. All statistical analyses were carried out with IBM SPSS Statistics, version 19 (IBM, Armonk, NY). The data are expressed as the mean (\pm s.d.) or percentage. BPV was measured as the s.d. of the clinic BP assessments, the awake ABP readings and the sleep ABP readings. The χ^2 -test was used to compare proportions. The independent samples *t* test was performed to test group differences in means. In the survival analyses for Hard CVD, duration of follow-up was defined as the months from the baseline to the first occurrence of a Hard CVD event, or last follow-up date of subjects who had Soft CVD or no CVD event; for All CVD events, the duration of follow-up was defined as the months from the baseline to the first occurrence of a Hard or Soft CVD event, or last follow-up date of subjects without any CVD event. All inferential statistics are based on the Cox regression analyses (see below) where the BPV measures are treated as continuous predictors. BP readings after the onset of the first CVD event were excluded from the computation of mean clinic BP and clinic BPV measures. Adjusted hazard ratios (HRs) with 95% confidence intervals (CIs) were based on univariate and multivariable Cox regression analyses. HRs for each BPV measure were expressed as HR per 5 mm Hg increase in BPV and per 1 s.d. increase in BPV. In preliminary analyses, we performed Cox regressions using all potential predictors except the BPV parameters. We included age, sex, body mass index, smoking status, the presence of diabetes, serum creatinine, cholesterol, and clinic BP at baseline. Those variables with $P < 0.10$ in the preliminary analysis were included as covariates in the primary analysis: the selected variables were age, diabetes, creatinine, smoking, and clinic SBP for All CVD, and age, diabetes, and creatinine for Hard CVD events. The null hypothesis concerning the effect of BPV on incident CVD was rejected when two-tailed $P < 0.05$. The results of a *post-hoc* power analysis for the Cox regression analyses are reported. In order to illustrate the univariate associations of BPV with incident CVD, we dichotomized the continuous predictors at their medians (the cutoff values were 13.3 mm Hg for clinic BPV and 12.2 mm Hg for sleep BPV) and present the separate Kaplan–Meier survival curves for those in the top and bottom halves of the BPV distribution.

RESULTS

Table 1 shows the baseline characteristics of subjects. The mean age was 67.0 ± 9.2 years; females outnumbered males; 44% had diabetes; and 56% of the subjects were on antihypertensive treatment.

During the follow-up period, 34 Hard CVD events and 58 All CVD events occurred. In univariate analyses of systolic BPV measures, BPV of sleep SBP was significantly associated with Hard CVD events, whereas BPV of clinic and both ambulatory BPV measures were significantly associated with All CVD events (**Table 2**). Although the results were similar for diastolic BPV when predicting Hard CVD events, only the BPV of clinic diastolic BP was a significant predictor of All CVD events (**Table 2**). The results of the multivariable Cox regression analyses for BPV that controlled for significant covariates are shown in **Table 3**. Ambulatory awake systolic

Table 1 | Baseline characteristics of subjects, N = 457

	Mean \pm s.d. or %	Range, min, max
Age, years	67.0 \pm 9.2	33, 88
Sex, male %	37.6	
Body mass index, kg/m ²	23.9 \pm 3.5	15.4, 46.3
Current smoking, %	24.9	
Type 2 diabetes, %	44.2	
Duration of hypertension, years	6.2 \pm 7.4	0, 54
Antihypertensive medications, %	55.6	
Total cholesterol, mg/dl	203 \pm 35	101, 306
Triglycerides, mg/dl	124 \pm 65	36, 400
Creatinine, mg/dl	0.77 \pm 0.21	0.30, 1.79
Clinic systolic BP, mm Hg	154 \pm 20	100, 210
Clinic diastolic BP, mm Hg	84 \pm 12	47, 116
24-h systolic BP, mm Hg	140 \pm 17	99, 199
24-h diastolic BP, mm Hg	79 \pm 10	55, 106
24-h pulse rate, bpm	68 \pm 9	42, 98
Awake systolic BP, mm Hg	146 \pm 18	102, 213
Awake diastolic BP, mm Hg	83 \pm 10	56, 110
Awake pulse rate, bpm	72 \pm 9	43, 103
Sleep systolic BP, mm Hg	129 \pm 19	85, 187
Sleep diastolic BP, mm Hg	73 \pm 10	47, 103
Sleep pulse rate, bpm	61 \pm 9	38, 93
Night/day ratio of SBP	0.89 \pm 0.09	0.60, 1.26
BPV (s.d.) of clinic systolic BP, mm Hg	13.7 \pm 4.6	2.9, 40.9
BPV (s.d.) of clinic diastolic BP, mm Hg	8.2 \pm 2.9	0, 33.2
BPV (s.d.) of Awake systolic BP, mm Hg	17.7 \pm 5.2	7.9, 36.4
BPV (s.d.) of Awake diastolic BP, mm Hg	10.5 \pm 2.5	5.1, 19.8
BPV (s.d.) of Sleep systolic BP, mm Hg	12.9 \pm 4.5	4.2, 30.5
BPV (s.d.) of Sleep diastolic BP, mm Hg	8.6 \pm 2.6	0, 21.1

BP, blood pressure; bpm, beats per minute; BPV, blood pressure variability; Max, maximum; Min, minimum; SBP, systolic BP.

BPV and diastolic BPV were not significantly associated with either outcome after adjustment for covariates. Therefore, for ambulatory BPV, only sleep systolic BPV and clinic systolic BPV were included in the models. For Hard CVD events, sleep systolic BPV was a significant predictor independent of clinic systolic BPV and the covariates; on the other hand, for All CVD events, clinic systolic BPV was a significant predictor independent of sleep systolic BPV and the other covariates (**Table 3**). The HR (95% CI) for clinic systolic BPV when predicting All CVD events did not change even when the use of antihypertensive medication(s) at baseline was entered into the model (HR, 1.48 per 5 mm Hg, 95% CI = 1.12–1.97, $P = 0.007$). Diastolic BPV exhibited a similar pattern of results to systolic BPV (shown in **Supplementary Table S1** online).

The Kaplan–Meier survival curves for those above and below the median of clinic systolic BPV (≥ 13.3 mm Hg or < 13.3 mm Hg for both Hard CVD and All CVD) are shown

Table 2 | Univariate analyses of systolic and diastolic BP variability for cardiovascular events

	Hazard ratio	Hard CVD		P	All CVD	
		95% CI	P		Hazard ratio	95% CI
<i>Systolic BP measures</i>						
BPV (s.d.) of clinic SBP per 5 mm Hg (per 1 s.d.)	0.95 (0.96)	0.61–1.48	0.82	1.58 (1.52)	1.25–2.00	<0.01
BPV (s.d.) of awake SBP per 5 mm Hg (per 1 s.d.)	1.28 (1.30)	0.96–1.72	0.09	1.27 (1.28)	1.02–1.59	0.04
BPV (s.d.) of sleep SBP per 5 mm Hg (per 1 s.d.)	1.58 (1.51)	1.15–2.17	<0.01	1.37 (1.33)	1.07–1.78	0.02
<i>Diastolic BP measures</i>						
BPV (s.d.) of clinic DBP per 5 mm Hg (per 1 s.d.)	1.10 (1.06)	0.58–2.08	0.77	1.55 (1.29)	1.12–2.14	<0.01
BPV (s.d.) of awake DBP per 5 mm Hg (per 1 s.d.)	1.42 (1.19)	0.75–2.68	0.28	1.29 (1.14)	0.79–2.11	0.31
BPV (s.d.) of sleep DBP per 5 mm Hg (per 1 s.d.)	2.04 (1.45)	1.14–3.65	0.02	1.59 (1.27)	0.99–2.57	0.06

These variables were entered one-by-one.

BP, blood pressure; BPV, BP variability; CI, confidence interval; CVD, cardiovascular disease; DBP, diastolic blood pressure; SBP, systolic BP.

Table 3 | Multivariable Cox regression analyses of SBP variability for cardiovascular events

	Hard CVD			All CVD		
	Hazard ratio	95% CI	P	Hazard ratio	95% CI	P
Age, mean per 10 years	2.27	1.37–3.76	<0.01	1.63	1.16–2.31	<0.01
Diabetes (yes or no)	2.99	1.35–6.63	<0.01	2.97	1.61–5.45	<0.01
Creatinine level per 0.1 mg/dl	1.24	1.05–1.46	0.01	1.12	0.98–1.28	0.09
Current smoking (yes or no)	—	—	—	1.61	0.88–2.94	0.12
Clinic SBP, mean per 10 mm Hg	—	—	—	1.18	1.03–1.35	0.02
BPV (s.d.) of clinic SBP, per 5 mm Hg (per 1 s.d.)	0.75 (0.76)	0.48–1.17	0.20	1.48 (1.44)	1.12–1.97	<0.01
BPV (s.d.) of sleep SBP, per 5 mm Hg (per 1 s.d.)	1.49 (1.43)	1.04–2.13	0.03	1.03 (1.02)	0.74–11.43	0.88

These variables were entered together for each endpoint.

BPV, blood pressure variability; CI, confidence interval; CVD, cardiovascular disease; SBP, systolic blood pressure.

in **Figure 1**. For Hard CVD events, the two survival curves are similar, whereas for All CVD, those with BPV of Clinic SBP >13.3 mm Hg had a significantly higher event rate than those with BPV <13.3 mm Hg. In terms of sleep systolic BPV, the survival curves were similar between the group with BPV of sleep systolic BPV \geq 12.2 mm Hg and the group with BPV <12.2 mm Hg for All CVD events, but was significantly different for Hard CVD events (**Figure 2**).

When the number of clinic BP assessments were divided into tertiles, those in the lowest tertile (1–22 readings), and middle tertile (23–50 readings) were at greater risk for Hard CVD than those in the highest tertile (51–78 readings) (HR = 3.91, 95% CI = 1.44–10.62, $P < 0.01$ and HR = 2.87, 95% CI = 1.07–7.73, $P = 0.04$, respectively); for All CVD, HR = 3.46, 95% CI = 1.60–7.50, $P < 0.01$ for lowest tertile and HR = 2.75, 95% CI = 1.35–5.59, $P < 0.01$ for middle tertile than those in the highest tertile. These are largely tautological, since clinic BP assessments obtained after a CVD event were not used. When we further added the interaction of tertile with clinic BPV to the Cox regression model, the interactions were not significant (all $P > 0.4$), indicating that the effect of clinic BPV on CVD risk was not associated with the number of clinic BP assessments. Therefore, there was no relationship between the number of BP readings and the predictability of events. Parallel analyses predicting all CVD yielded equivalent results.

Although there were significant correlations of the clinic systolic BPV with the ambulatory sleep and awake systolic BPV measures ($r = 0.19$, $P < 0.01$; $r = 0.20$, $P < 0.01$, respectively), multicollinearity among these measures was clearly not an issue (variance inflation factor <2.0).

Post-hoc power calculations for our study showed that the analyses predicting Hard CVD events had 80% power to detect a HR of 1.62 (per 1 s.d.) and 90% power to detect a HR of 1.76 (per 1 s.d.). The analyses predicting All CVD events had 80% power to detect a HR of 1.45 (per 1 s.d.) and 90% power to detect a HR of 1.54 (per 1 s.d.).

DISCUSSION

To the best of our knowledge, this is the first clinical study to compare the associations of clinic BPV and ambulatory BPV in the same database. The results showed that clinic BPV was a significant predictor for All CVD events, whereas sleep BPV was associated with Hard CVD events. Thus, each BPV measure is separately useful for the assessment of future CVD events.

Clinic BPV and CV events

In the present study, clinic BPV was shown to be associated with all CVD events, which included both Hard and Soft CVD events, but was not associated with Hard CVD events alone.

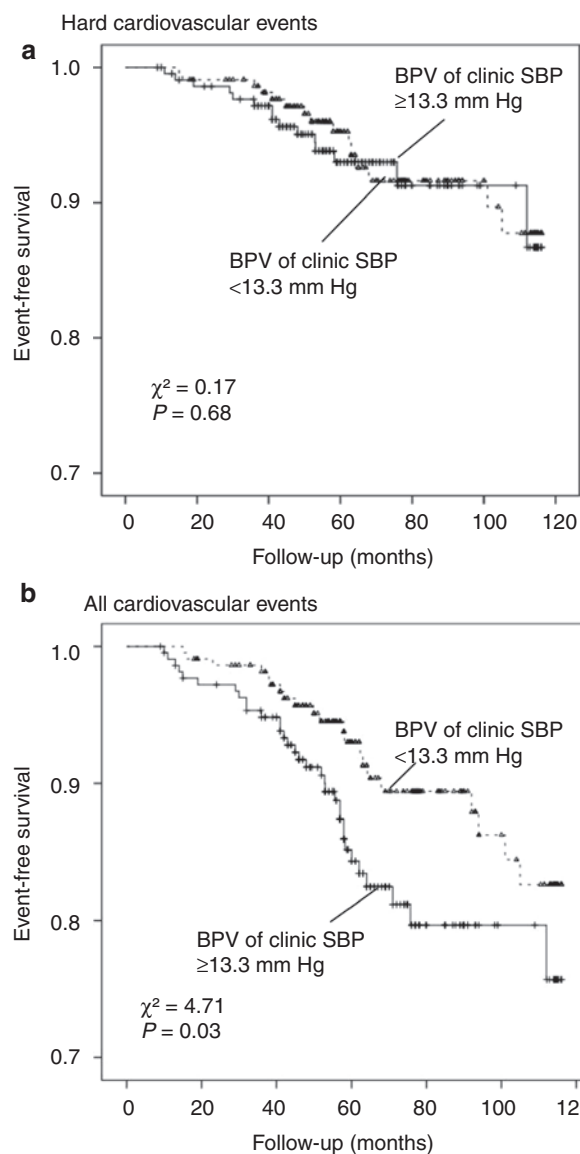


Figure 1 | Kaplan–Meier curves of event-free survival for two categories of clinic systolic blood pressure variability (BPV) (s.d. ≥ 13.3 vs. < 13.3 mm Hg) for (a) Hard cardiovascular events and (b) All cardiovascular events. SBP, systolic blood pressure.

In previous studies, clinic BP was recorded every 4 months,³² or only two to three times³³ for calculating clinic BPV. In this study, clinic BP was recorded every month for most of the subjects. Because there are seasonal variations in BP, it would be natural to have a greater BPV if the measurement interval was very long (4 months or more). Our data are therefore more typical of real clinical practice than previous studies because BP medications are commonly titrated once or twice every month until the BP is stabilized. Variations in BP at clinic visits reflect many factors: the emotional state, position, respiratory cycle, diet, salt intake, alcohol ingestion, physical activity, and amount of rest of the subject, as well as the time of day and room temperature during the measurement, and the potential presence of other nonstandardized conditions for BP measurements.^{34,35} The association between clinic BPV and

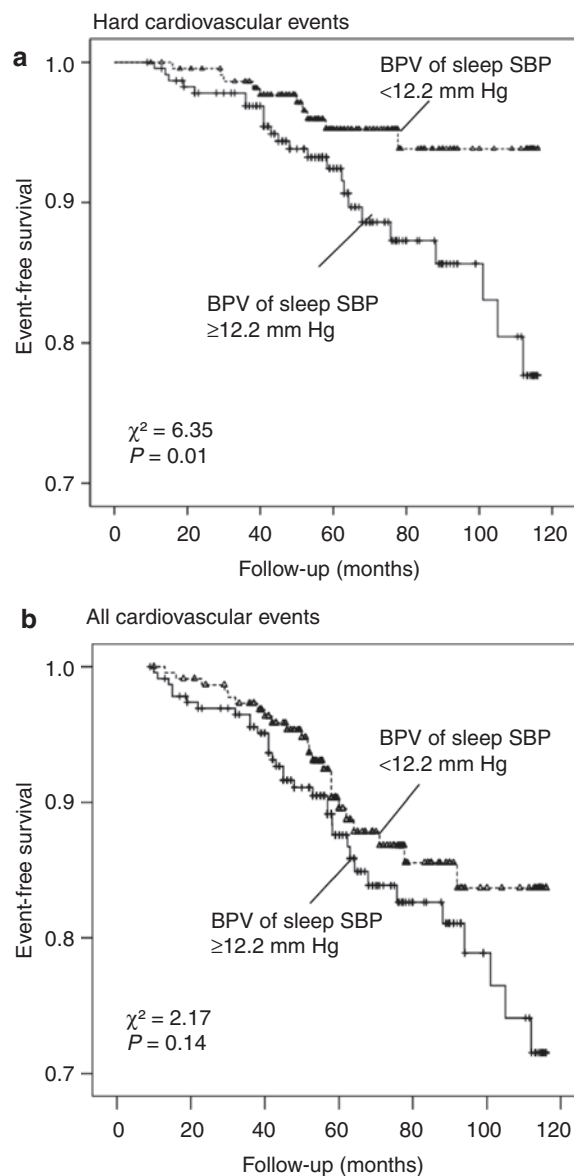


Figure 2 | Kaplan–Meier curves of event-free survival for two categories of sleep systolic blood pressure variability (BPV) (s.d. ≥ 12.2 vs. < 12.2 mm Hg) by ambulatory blood pressure monitoring for (a) Hard cardiovascular events and (b) All cardiovascular events. SBP, systolic blood pressure.

CVD prognosis may be especially relevant to high-risk populations. In this study, almost half of the subjects had type 2 diabetes, putting them at high risk for CVD. In the Framingham study, clinic BP lability was not associated with cardiovascular events.³⁶ A more recent study also failed to show a positive relationship between ambulatory BPV and CV events.²⁰ In our series, although the clinic BP was measured under standard conditions to the greatest extent possible, patients with advanced atherosclerosis could have had greater variability. We speculate that patients with greater clinic BPV are more likely to have cardiovascular target organ damage, which would make them more susceptible to all CVD events, including congestive heart failure, angina, and transient ischemic attacks.^{23,32,33}

ABP variability and CV events

In this study, ABP variability, especially BPV during sleep, was associated with Hard CVD events. This provides further confirmation of the previously reported utility of ABPM in clinical practice.^{4,5,17,21} This is in line with our previous subgroup analysis of diabetic subjects in this database¹¹ and with another Japanese study.³⁷ The clinical significance of ambulatory BPV has also been shown in previous studies.^{4–5,10} Increased nighttime systolic BPV has been shown to be an independent risk factor for stroke in subjects with isolated systolic hypertension,² and untreated essential hypertension,¹⁰ both of which are consistent with the present study. Ambulatory BPV is influenced by various daily activities, such as diet, exercise, rest, change in temperature, sleep, and mental stress, and it reflects dynamic changes of BP during daily life. On the other hand, because clinic BP is measured under relatively controlled conditions, the mechanism of fluctuation is likely to be completely different. It can be speculated that hemodynamic instability under ambulatory conditions reflects more advanced atherosclerosis and further fluctuations of BP than those in clinic BPV. Therefore, we conclude that ABPM is an important tool for the prevention of Hard CVD events.

There are several limitations in this study. The sample is a heterogeneous mixture of newly diagnosed hypertension patients, hypertension patients receiving nonpharmacological treatments, and white-coat hypertension patients. However, this type of heterogeneity is the norm for general internal medicine clinics. Second, the sample was only intermediate in size. The other limitations are the relatively large number of covariates for the event number. Clinic BPV during follow-up, which is a marker of longitudinal BP change, and ambulatory BPV at baseline, which is a marker of BP fluctuations during one day of daily life, are completely different in nature; therefore, while one may be a better predictor of a specific outcome, it would be problematic to conclude that one is inherently superior to the other based on this or any other study. Finally, given that the present study only had 95% power to detect standardized HRs of ~1.9 (or larger), and associations of less than this magnitude would certainly be of interest, readers should not interpret the nonstatistically significant associations in this study as evidence that those associations are not present in the larger population.

In hypertensive patients, increased visit-to-visit clinic BPV was associated with increased incidence of all CVD events, and increased sleep BPV was associated with increased incidence of Hard CVD events. Clinic BPV has the advantage of providing data without any special device, but requires time (multiple visits) to collect sufficient readings to calculate BPV, even though the number of clinic BP readings was not predictive of CVD risk. On the other hand, the numerous BP readings obtained from a single 24-h ABPM allow the calculation of ambulatory BPV in daily life, but the generalizability of the result beyond the single day being measured is unclear, and ABPM obviously requires special equipment. Each BPV measure has advantages and disadvantages, and can be separately used as a marker of future cardiovascular outcomes.

Supplementary material is linked to the online version of the paper at <http://www.nature.com/ajh>

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